

# *Inspection and maintenance decisions based on imperfect inspections*

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## Presentation outline

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- Introduction
- Model concept + gamma process
- Perfect & imperfect inspections
- Case study: inspecting a hydrogen dryer
- Conclusions

# *Introduction*

## Research objective

“To develop an easy-to-use model for making optimal inspection and maintenance decisions for pressurized vessels in the process industry, under the assumption that field measurement data is imperfect.”



*Applying a stochastic process to the corrosion  
degradation mechanism*

## Corrosion state functions

The thinning due to corrosion state function is given by (American Petroleum Institute, 2000):

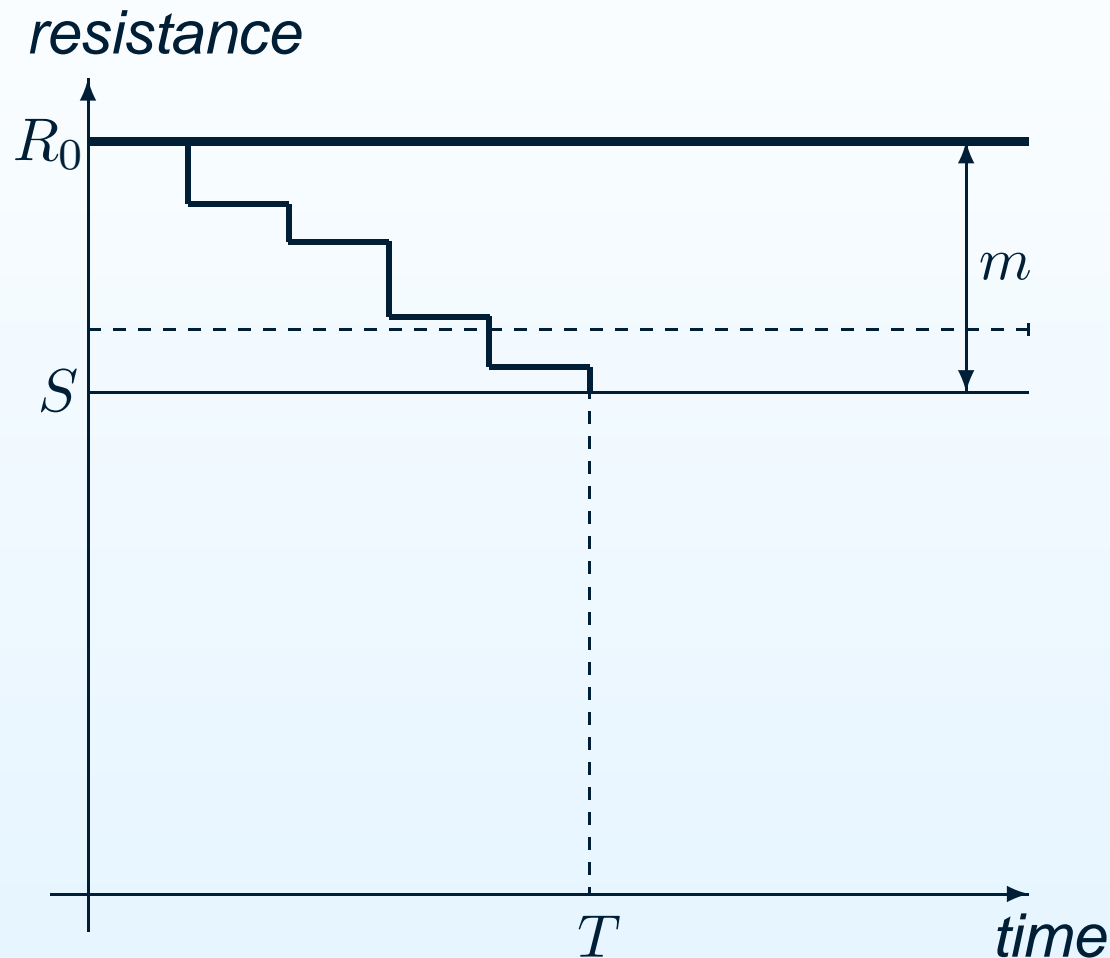
$$g(t) = \underbrace{S \left( 1 - \frac{C \times t}{th} \right)}_{Resistance} - \underbrace{\left( \frac{P \times d}{2th} \right)}_{Stress},$$

$S =$	Material strength [MPa = 10bar]	$t =$	Time [yr]
$C =$	Corrosion rate [mm/yr]	$d =$	Diameter [mm]
$P =$	Operating pressure [bar]	$th =$	Thickness [mm]

The limit state is given by:

$$g(t) = 0 \Leftrightarrow th - C \times t = \frac{P \times d}{2S}$$

# Stochastic deterioration



← Initial thickness

← Corrosion allowance

← Failure

The safety margin  
is given by:

$$m = th - \frac{P \times d}{2S}$$

*Bayesian gamma stochastic process decision model*



## Gamma process for corrosion

- The corrosion is modelled as a gamma process with independent increments:  $X(t) \sim \text{Ga}(x|\alpha t, \beta)$  with  $t \geq 0$ .
- Assuming linear degradation we want

$$\mathbb{E}(X(t)) = \mu t \text{ and } \text{Var}(X(t)) = \sigma^2 t,$$

where  $\mu$  is the average amount of deterioration per time unit and  $\sigma$  is the standard deviation of the deterioration.

- For  $X(t)$  it holds that  $\mathbb{E}(X(t)) = \frac{\alpha}{\beta}t$  and  $\text{Var}(X(t)) = \frac{\alpha}{\beta^2}t$ , therefore the parameters of the process are given by

$$\mu = \frac{\alpha}{\beta}, \sigma^2 = \frac{\alpha}{\beta^2} \quad \Rightarrow \quad \alpha = \left(\frac{\mu}{\sigma}\right)^2, \beta = \frac{\mu}{\sigma^2}.$$

## Fixing the standard deviation

- The mean and variance are uncertain, but by fixing the standard deviation relative to the mean, i.e.

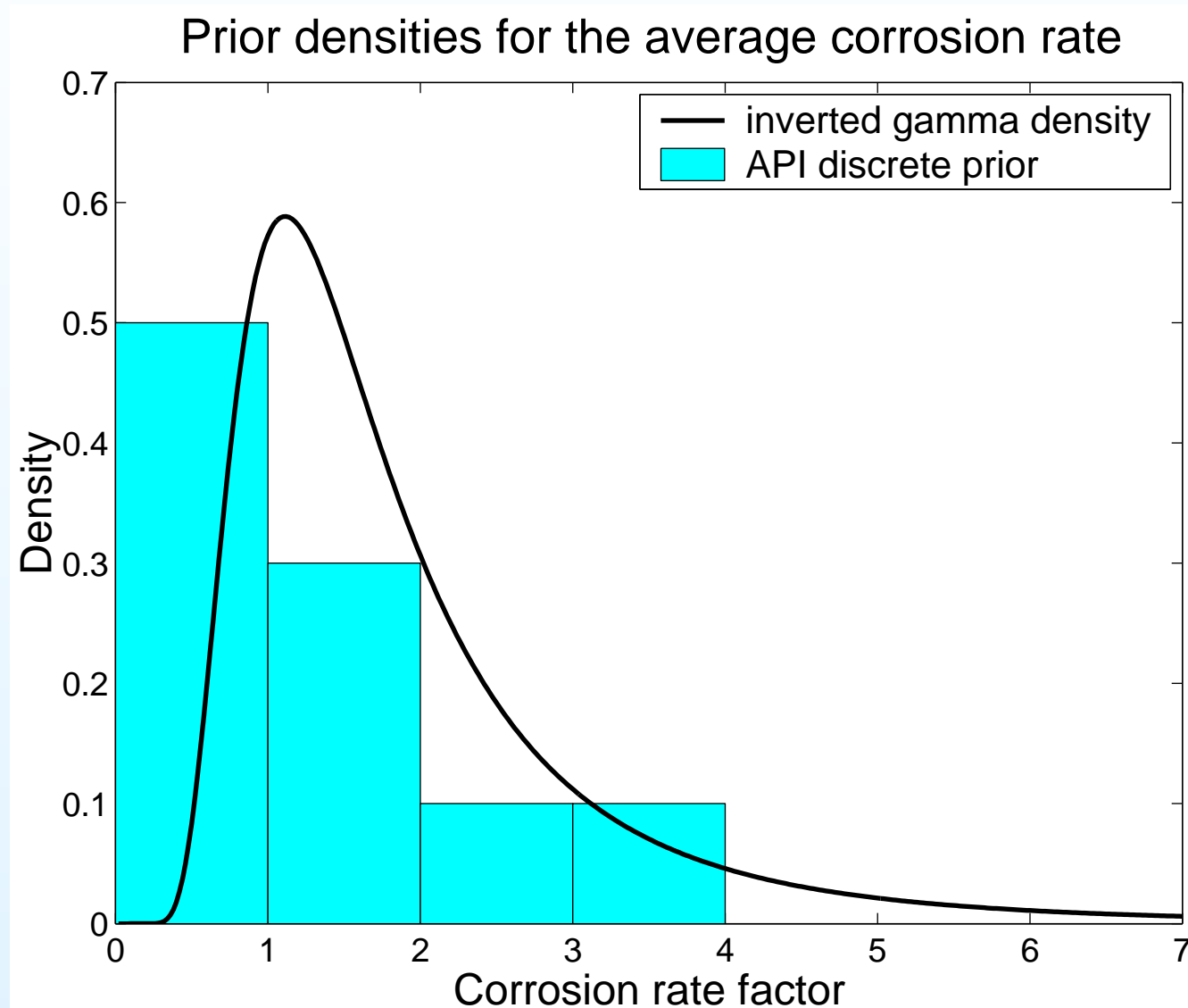
$$\sigma = \text{COV} \times \mu,$$

results in ( $\nu = \text{COV}$ )

$$X(t) \sim \text{Ga} \left( x \mid \frac{t}{\nu^2}, \frac{1}{\mu\nu^2} \right), \quad \nu > 0.$$

- This manipulation avoids the need for an assessment of the COV in the absence of sufficient data.

## Inverted gamma prior



## Perfect inspections (1)

- Prior density  $\pi(\mu) = \text{lg}(\mu|a, b)$ , where the inverted gamma density with parameters  $a > 0$  and  $b > 0$  is defined as:

$$\text{lg}(x|a, b) = \frac{b^a}{\Gamma(a)} \left(\frac{1}{x}\right)^{a+1} \exp\left\{-\frac{b}{x}\right\}, \quad x \geq 0.$$

- Posterior density with one inspection:

$$\pi(\mu|x) = \text{lg}\left(\mu \left| \frac{t}{\nu^2} + a, \frac{x}{\nu^2} + b\right.\right) \propto l(x|\mu)\pi(\mu),$$

where  $l(x|\mu)$  is the likelihood of measurement  $x$  ( $x > 0$ ) at time  $t$  and  $\pi(\mu)$  is the prior for the average corrosion rate.

## Perfect inspections (2)

- Multiple perfect inspections:

$$\begin{aligned}\pi(\mu|x_1, \dots, x_n) &= \\ &= \text{lg} \left( \mu \left| \frac{\sum_{i=1}^n t_i - t_{i-1}}{\nu^2} + a, \frac{\sum_{i=1}^n x_i - x_{i-1}}{\nu^2} + b \right. \right)\end{aligned}$$

- due to our choice of  $\nu = \sigma/\mu$ , the posterior for multiple perfect inspections only depends on the last measurement  $x_n$  at time  $t_n$ . We assume that  $t_0 = x_0 = 0$ .

## Imperfect inspections (1)

- Current non-destructive testing (NDT) techniques, like ultrasonic thickness measurements, are not capable of measuring the exact material thickness.
- To extend the existing Bayesian updating model such that it can cope with imperfect inspections, we use a stochastic process  $Y(t)$  consisting of the actual deterioration process  $X(t)$  plus a normally distributed measurement error:

$$Y(t) = X(t) + \epsilon \quad \text{with} \quad \epsilon \sim \mathcal{N}(0, \sigma_\epsilon).$$

## Imperfect inspections (2)

- likelihood for 1 inspection:

$$l(y|\mu) = f_{Y(t)}(y) = f_{X(t)+\epsilon}(y) = \int_{-\infty}^{\infty} f_{X(T)}(y - \epsilon) f_{\epsilon}(\epsilon) d\epsilon$$

- likelihood for  $K > 1$  inspections:

$$\begin{aligned} l(y_1, \dots, y_K | \mu) &= \prod_k l_{Y(k)-Y(k-1)}(y_k - y_{k-1} | \mu) = \\ &= \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} \prod_k f_{D_k}(y_k - y_{k-1} - \delta_k) f(\delta_1, \dots, \delta_k) d\delta_1 \cdots d\delta_k \end{aligned}$$

where  $D_k = Y(k) - Y(k - 1)$  and  $\delta_k = \epsilon_k - \epsilon_{k-1}$ .

## Imperfect inspections (3)

- Using Monte Carlo integration, the likelihood is approximated by

$$l(y_1, \dots, y_K | \mu) \approx \frac{1}{N} \sum_{j=1}^N \left[ \prod_k \text{Ga} \left( d_k - \delta_k^{(j)} \mid \frac{t_k - t_{k-1}}{\nu^2}, \frac{1}{\mu \nu^2} \right) I_{[0, \infty)}(d_k - \delta_k^{(j)}) \right],$$

as  $N \rightarrow \infty$  and where  $d_k = y_k - y_{k-1}$  and  $\delta_k^{(j)} = \epsilon_k^{(j)} - \epsilon_{k-1}^{(j)}$ .



## Expected average costs per unit time

- A useful cost based decision criterium is the expected average costs per time unit. For each inspection interval  $\Delta k$  we calculate the ratio of the expected (cumulative) cycle costs over the expected cycle length:

$$C(\rho, \Delta k) = \frac{\sum_{i=1}^{\infty} c_i(\rho, \Delta k) p_i(\rho, \Delta k)}{\sum_{i=1}^{\infty} i p_i(\rho, \Delta k)},$$

where  $\rho$  defines the replacement level (i.e. the corrosion allowance)

*Case study: inspecting a hydrogen dryer*

## Vertical drum input data

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Service start:	1977	
Operating pressure:	32	bar
Drum diameter:	1180	mm
Initial thickness:	15+12% = 16.8	mm
Corrosion allowance:	4.5	mm

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### *Ultrasonic thickness measurements:*

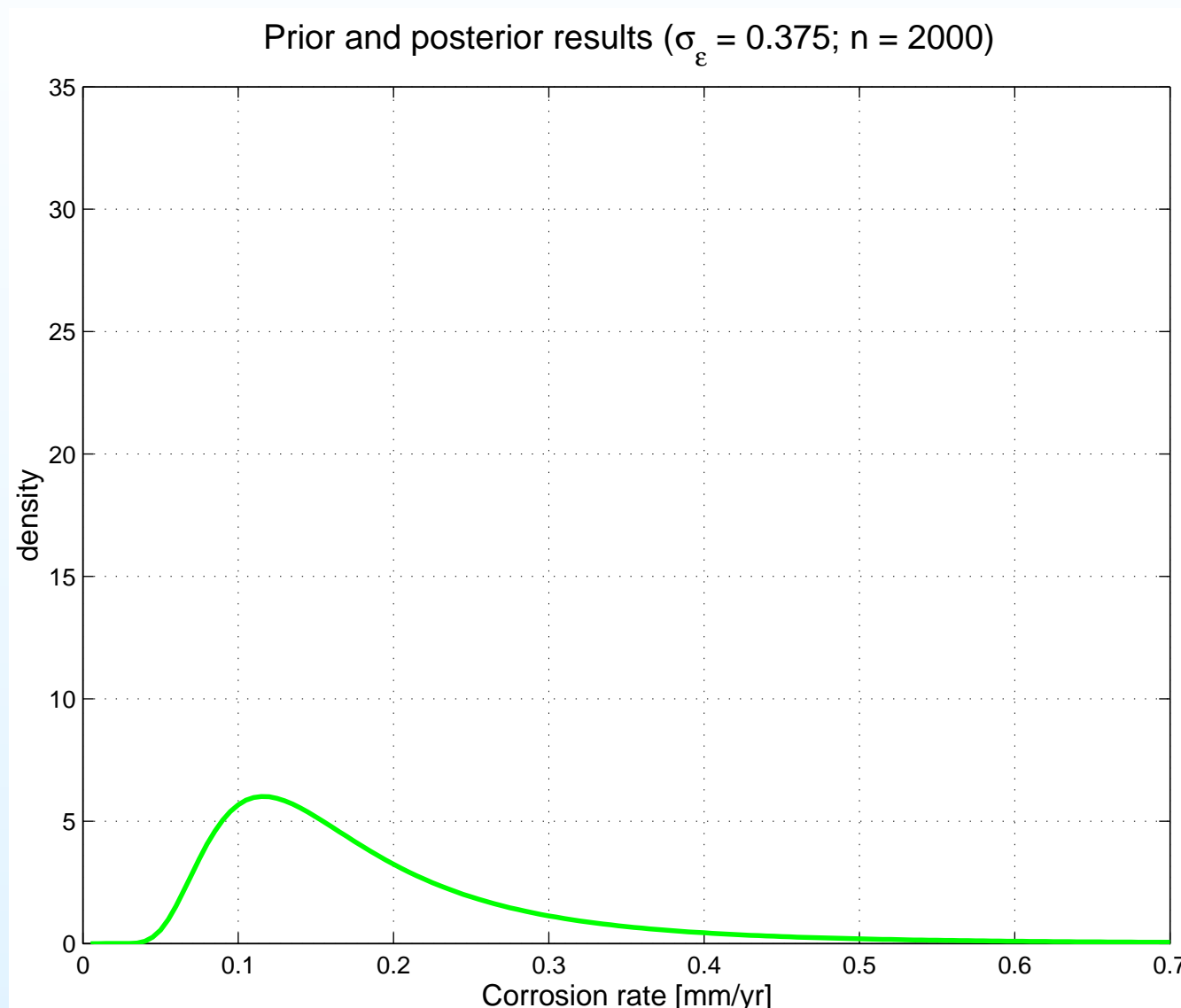
1986:	15.6	mm
1990:	14.6	mm
1994:	14.2	mm
1998:	13.8	mm

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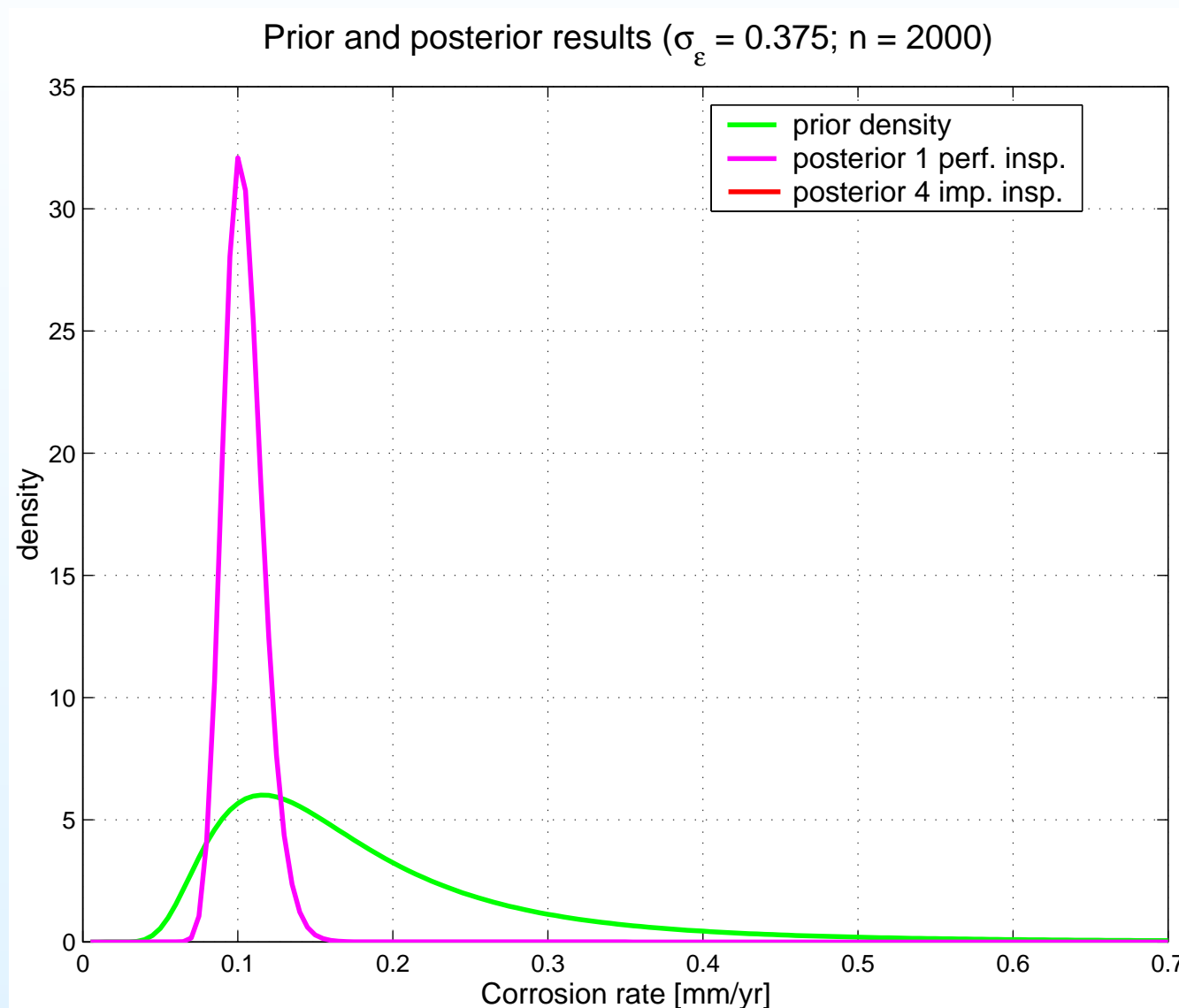
### *Costs:*

Inspection:	10,000	\$
Preventive replacement:	50,000	\$
Corrective replacement:	1,000,000	\$

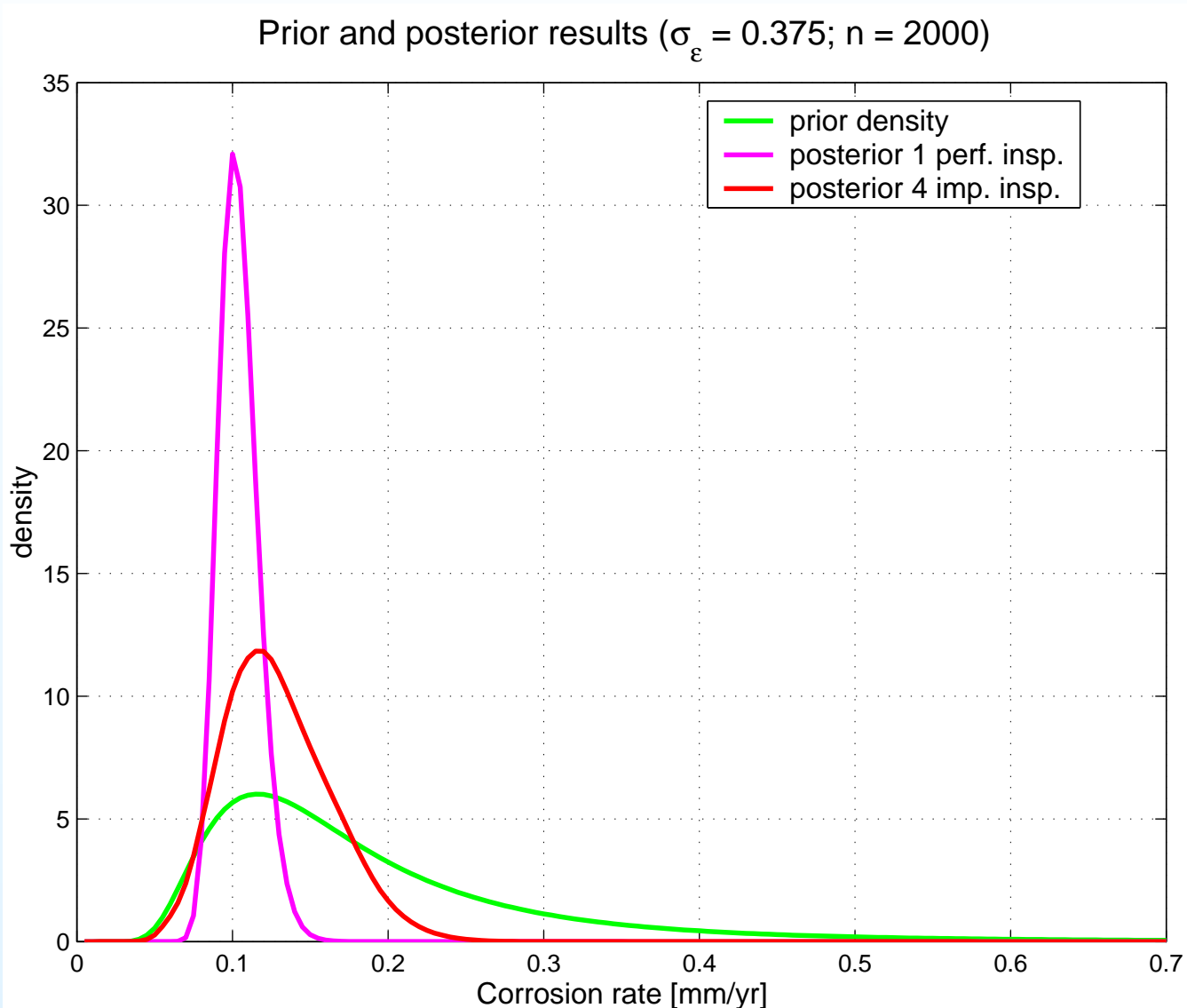
# Continuous prior and posterior



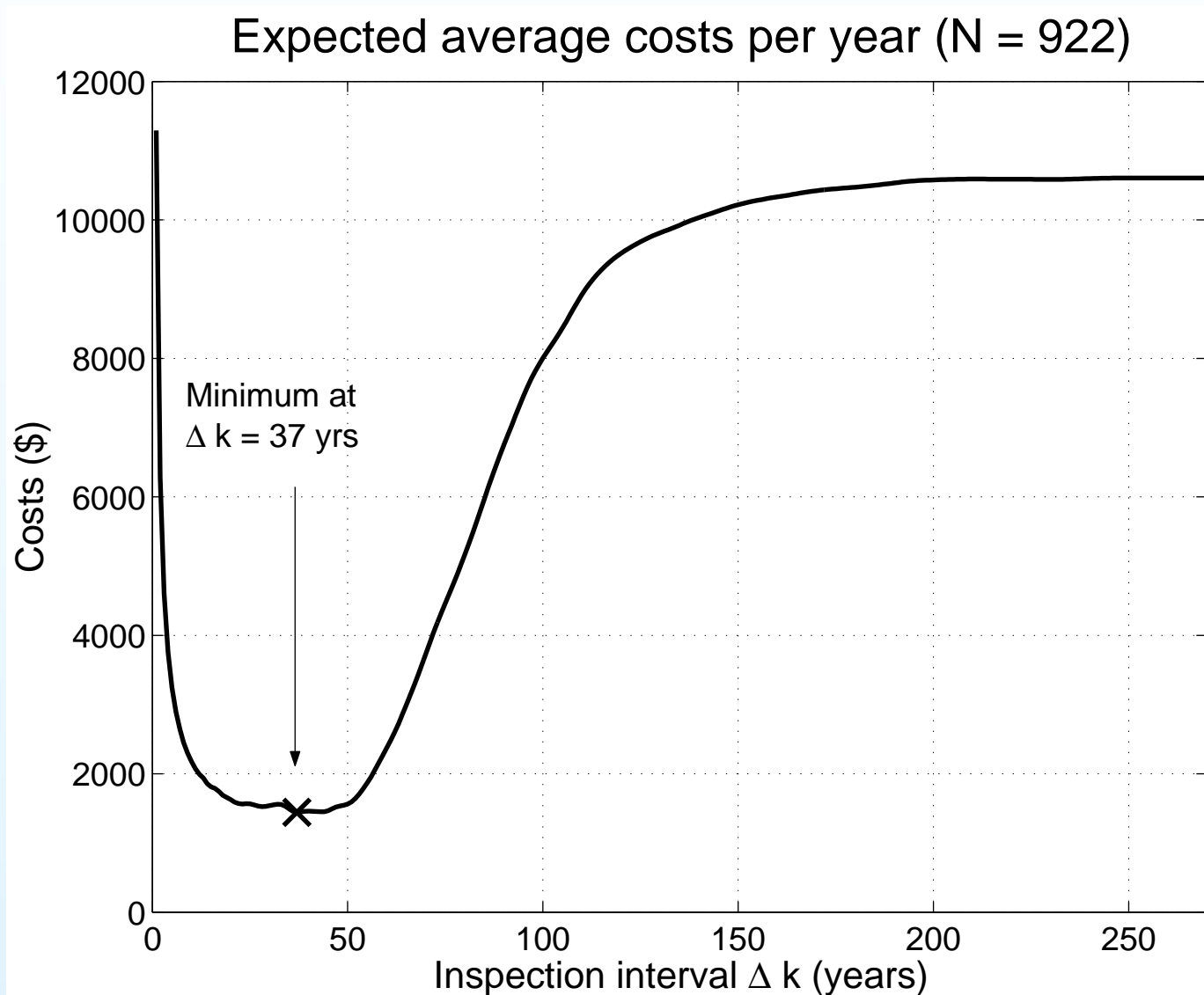
# Continuous prior and posterior



# Continuous prior and posterior



# Expected average costs per unit time



## Conclusions

- The Bayesian stochastic process with gamma distributed increments is a very good alternative to current methodologies. The expected average costs per year cost criterium is well suited for safe and economical decisions.
- Computationally this model is inefficient, but with the right assumptions and a proper implementation this can be greatly improved:
  - For example: we can choose an isotropic time grid (i.e.  $t = 1/\alpha = \text{COV}^2$ ), such that the amounts of deterioration are exponentially distributed. This enables more analytic calculations, reducing the computational load.